
Psychometric characteristics of the Italian version of the revised Fibromyalgia Impact Questionnaire using classical test theory and Rasch analysis

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ABSTRACT

Objectives. The aim of this study was to perform a psychometric analysis of the Italian Fibromyalgia Impact Questionnaire Revised version (FIQR), using both classical test theory (CTT) and Rasch analysis (RA) in order to better analyse its construct validity and provide a rational basis for a possible improvement of its metric quality.

Methods. The study involved 503 patients with fibromyalgia (FM) (423 women and 80 men) with mean age of 51.3 ± 10.1 years (range 19–74) and mean duration of symptoms of 11.1 ± 8.7 years (range 1–30). All patients completed the Italian FIQR during their clinical visit. The translation and cultural adaptation process of the Italian FIQR followed the published guidelines and no local adjustments were made except for a slight adaptation of item 13 related to "energy".

Results. Factor Analysis revealed two salient dimensions: function (items 1–9) and symptoms (items 12–21). RA was thus performed on these two subscales. Rating scale diagnostics suggested collapsing the eleven rating categories of the scale into five. After combining these rating categories, RA showed that most items of each of the two subscales fitted the respective constructs to measure (MnSq 0.7–1.3). The reliability levels of the two subscales were higher than 0.80.

Conclusion. This study provides psychometric evidence of the reliability, internal validity and two-dimensional structure of the FIQR in a FM population. Our results support the use of two separate subscales for "function" and "symptoms", and provide a useful starting point for further refinement of the scale.

Introduction

Fibromyalgia (FM) is a chronic multi-symptom disease (1-3) affecting ap-

proximately 2–3% of the general population (more than 90% of patients are female), with pain as possibly its most important symptom (1, 4, 5). It consequently tends to have a profound impact on health-related quality of life (HRQL) (6) and is associated with high rates of healthcare resource use and an increased risk of inability to work (7), suicide and suicide attempts (8). Patients with FM report disability in daily living activities that is as severe as that reported by patients with rheumatoid arthritis (6, 9), and more severe than that reported by patients with osteoarthritis (10) or other painful conditions (11-14).

It has been shown that measuring HRQL is a key area of assessment when screening for disability and improving communication between patients and clinicians, and it is considered to be essential by regulatory agencies when contemplating the approval of medications for chronic pain status (15, 16).

The instruments used to measure physical function and health status are generally divided into generic and specific measures (17, 18): while the former provide a broad picture of health status over a range of conditions, the latter are more sensitive to the specific disorder and therefore more likely to reflect clinically important changes (18).

The Fibromyalgia Impact Questionnaire (FIQ) (19) and its revised version (FIQR) (20) are the most widely used FM-specific measures to assess the full spectrum of problems related to FM and its response to therapy. The FIQ was in the past recommended as a primary efficacy endpoint measure in FM clinical trials (21), and long considered to be the standard measure of multi-dimensional function/HRQL in FM patients. However, since questions were raised about its scoring (response options) and content validity (absence

of questions related to FM symptoms such as cognitive dysfunction, tenderness, balance problems, and environmental sensitivity), a revised version (FIQR) was developed to address these limitations without losing any of the essential properties of the original (22). The FIQR is currently the recommended means for assessing function and HRQL in FM patients (23). The psychometric properties of FIQR have been tested using only the classical test theory approach and focusing only on basic statistics, reliability item analysis, Cronbach's alpha and external construct validity (22, 24, 25). To our knowledge, no study has analysed the FIQR using Rasch analysis, a statistical method for testing if the properties of a questionnaire comply with a wide range of psychometric requirements that cannot be analysed using classical test theory.

The aim of this study was to perform a comprehensive psychometric analysis of the Italian version of the FIQR in a large sample of FM patients, using both classical test theory (CTT) and Rasch analysis, in order to examine its main metric properties (*i.e.* rating scale functioning, internal construct validity, reliability indices, dimensionality) and provide insights for a possible improvement of the questionnaire.

Patients and methods

Study population

The study was carried out in 503 FM patients (423 women and 80 men), mean age 51.3±10.1 years (range 19–74) and mean duration of symptoms 11.1±8.7 years (range 1–30). All patients were recruited from the Department of Rheumatology of the Politechnic University of the Marche, Ancona, Italy, and were considered eligible if they were at least 18 years old and met the criteria for fibromyalgia defined by the American College of Rheumatology (ACR) (26). The main exclusion criteria were: an unstable medical or psychiatric illness or current primary psychiatric diagnosis including severe depression, pain due to traumatic injury or structural/regional rheumatic disease, rheumatoid arthritis, inflammatory arthritis, and autoimmune disease.

Patients were examined for FM symptoms and clinical findings in order to confirm the diagnosis, and then they were asked to complete the Italian FIQR.

The disease-related symptoms were also evaluated by the Fibromyalgia Assessment Status (FAS) (27). Briefly, the FAS is a short, self-administered

VERSIONE ITALIANA DEL REVISED FIBROMYALGIA IMPACT QUESTIONNAIRE (ITALIAN-FIQR)

Nome: _____ Et : _____
 Cognome: _____ Durata dei sintomi (anni): _____

DOMINIO 1 – FUNZIONE FISICA

Per ognuna delle seguenti NOVE domande, segni la casella che meglio indica il grado di difficolt  da Lei avvertito, *nel corso dell'ultima settimana*, nel compiere ciascuna delle attivit  elencate, a causa della fibromialgia: (scegliere un solo numero).

1. SPAZZOLARE O PETTINARE I CAPELLI

Nessuna difficolt  0 1 2 3 4 5 6 7 8 9 10 Estrema difficolt 

2. CAMMINARE ININTERROTTAMENTE PER 20 MINUTI

Nessuna difficolt  0 1 2 3 4 5 6 7 8 9 10 Estrema difficolt 

3. PREPARARE I PASTI

Nessuna difficolt  0 1 2 3 4 5 6 7 8 9 10 Estrema difficolt 

4. PASSARE L'ASPIRAPOLVERE E LAVARE I PAVIMENTI

Nessuna difficolt  0 1 2 3 4 5 6 7 8 9 10 Estrema difficolt 

5. SOLLEVARE E PORTARE LE BORSE DELLA SPESA

Nessuna difficolt  0 1 2 3 4 5 6 7 8 9 10 Estrema difficolt 

6. SALIRE UN PIANO DI SCALE

Nessuna difficolt  0 1 2 3 4 5 6 7 8 9 10 Estrema difficolt 

7. CAMBIARE LE LENZUOLA DEL LETTO

Nessuna difficolt  0 1 2 3 4 5 6 7 8 9 10 Estrema difficolt 

8. STARE SEDUTO SU UNA SEDIA PER ALMENO 45 MINUTI

Nessuna difficolt  0 1 2 3 4 5 6 7 8 9 10 Estrema difficolt 

9. ANDARE A FARE LA SPESA

Nessuna difficolt  0 1 2 3 4 5 6 7 8 9 10 Estrema difficolt 

DOMINIO 1 – SUBTOTALE: _____

DOMINIO 2 – STATO DI SALUTE GENERALE

Per ognuna delle seguenti DUE domande, segni la casella che meglio indica il grado di difficolt  da Lei avvertito, *nel corso dell'ultima settimana*, nel compiere ciascuna delle attivit  elencate, a causa della fibromialgia: (scegliere un solo numero).

10. LA FIBROMIALGIA MI HA IMPEDITO DI PORTARE A TERMINE I LAVORI/COMPITI DELLA SETTIMANA

Mai 0 1 2 3 4 5 6 7 8 9 10 Sempre

11. SONO STATO/A COMPLETAMENTE SOPRAFFATTO/A DAI SINTOMI DELLA FIBROMIALGIA

Mai 0 1 2 3 4 5 6 7 8 9 10 Sempre

DOMINIO 2 – SUBTOTALE: _____

DOMINIO 3 – SINTOMI

Per ognuna delle seguenti DIECI domande, segni la casella che meglio indica il grado di severità del sintomo da Lei avvertito, nel corso dell'ultima settimana, a causa della fibromialgia: (scegliere un solo numero).

| | | | | | | | | | | | |
|--|--------------------------|--------------------------|--------------------------|--------------------------|--------------------------|--------------------------|--------------------------|--------------------------|--------------------------|--------------------------|------------------------------------|
| 12. ASSEGNI UN PUNTEGGIO AL SUO LIVELLO DI DOLORE | | | | | | | | | | | |
| Nessun dolore | <input type="checkbox"/> | Estremo dolore |
| | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 13. ASSEGNI UN PUNTEGGIO AL SUO LIVELLO DI STANCHEZZA | | | | | | | | | | | |
| Nessuna stanchezza | <input type="checkbox"/> | Estrema stanchezza |
| | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 14. ASSEGNI UN PUNTEGGIO AL SUO LIVELLO DI RIGIDITA' | | | | | | | | | | | |
| Nessuna rigidità | <input type="checkbox"/> | Estrema rigidità |
| | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 15. ASSEGNI UN PUNTEGGIO ALLA QUALITA' DEL SUO SONNO | | | | | | | | | | | |
| Ben riposato/a al risveglio | <input type="checkbox"/> | Estremamente stanco/a al risveglio |
| | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 16. ASSEGNI UN PUNTEGGIO AL SUO GRADO DI DEPRESSIONE | | | | | | | | | | | |
| Per nulla depresso/a | <input type="checkbox"/> | Estremamente depresso/a |
| | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 17. ASSEGNI UN PUNTEGGIO AI SUOI PROBLEMI DI MEMORIA | | | | | | | | | | | |
| Buona memoria | <input type="checkbox"/> | Scarsissima memoria |
| | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 18. ASSEGNI UN PUNTEGGIO AL SUO GRADO DI ANSIA | | | | | | | | | | | |
| Per nulla ansioso/a | <input type="checkbox"/> | Estremamente ansioso/a |
| | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 19. ASSEGNI UN PUNTEGGIO AL SUO LIVELLO DI DOLORABILITA' (dolore percepito al tatto) | | | | | | | | | | | |
| Nessuna dolorabilità | <input type="checkbox"/> | Estrema dolorabilità |
| | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 20. ASSEGNI UN PUNTEGGIO AI SUOI DISTURBI DI EQUILIBRIO | | | | | | | | | | | |
| Nessun problema di equilibrio | <input type="checkbox"/> | Scarsissimo equilibrio |
| | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 21. ASSEGNI UN PUNTEGGIO AL LIVELLO DELLE SUA SENSIBILITA' AI RUMORI, ALLE LUCI, AGLI ODORI ED AL FREDDO | | | | | | | | | | | |
| Nessuna sensibilità | <input type="checkbox"/> | Estrema sensibilità |
| | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |

DOMINIO 3 – SUBTOTALE: _____

PUNTEGGIO:

1. Sommare i punteggi di ogni singolo item per i tre domini (funzione fisica, stato di salute generale e sintomi);
2. Dividere per 3 il punteggio relativo al dominio FUNZIONE FISICA, lasciare il punteggio inalterato per il dominio STATO GENERALE DI SALUTE e dividere per 2 il punteggio del dominio SINTOMI
3. Sommare i risultanti punteggi dei 3 domini per ottenere il totale

PUNTEGGIO TOTALE FIQR

Fig. 1. Italian version of the Fibromyalgia Impact Questionnaire Revised version (FIQR).

Translation and cultural adaptation of the FIQR

The Fibromyalgia Impact Questionnaire Revised version (FIQR) is an updated version of the FIQ (19) that was developed by Bennett *et al.* (20) to address the limitations of the original FIQ (22). The new version has 21 items (all based on an 11-point numeric rating scale of 0–10, with higher scores indicating greater disease impact), and covers the three domains of function, overall impact, and symptoms. FIQR is an improvement on the former scale mainly due to the addition of new questions related to memory, tenderness, balance, and environmental sensitivity. As in the FIQ, all questions refer to the previous seven days, and the total maximum score is 100, with higher scores indicating greater disease impact (20). The FIQR score is the sum of the three domain scores: the summed score for the 9-item function domain (range 0–90) is divided by three; the summed score for the 2-item overall impact domain (range 0–20) remains as it is; and the summed score for the 10-item symptom domain (range 0–100) is divided by two.

The translation and cultural adaptation process followed established guidelines (29). The original English version of the FIQR was translated into Italian by two independent professional translators (Fig. 1). One translator was aware of the questionnaire, the other was not. The two forward translations were reviewed and compared with one another and with the original English version by the translators and one of the authors. Subsequently, a pilot test was run including a cognitive debriefing (*i.e.* an interview by a psychologist about the clarity, intelligibility, appropriateness, and cultural relevance of the target language version), and the final adaptation was prepared by experts. No local adjustments were made, except for a slight adaptation of item 13 related to “energy”, in which the descriptors are “0=no fatigue” and “10=extreme fatigue”. This was because the word “fatigue” is better received by Italian patients, and it also represents one of the major domains classified as important by patients (30).

index combining questions related to non-articular pain (range 0–10), fatigue (range 0–10) and quality of sleep (range 0–10), that provides a single composite measure of disease activity. The total score is calculated by summing the three subscores and dividing

the result by three (range 0–10). The FAS has demonstrated validity, reliability and responsiveness (27, 28). Our local Human Ethics Committees approved the research protocol and all patients gave their written informed consent.

Statistical analysis

We combined the approaches of CTT and Rasch analysis in order to investigate the psychometric properties of FIQR. In accordance with CCT, the internal consistency of the FIQR was assessed using Cronbach's coefficient alpha (values > 0.85–0.90 are desirable for individual judgments), and item-to-total correlation (the usual rule of thumb is that an item should correlate with the total score with $r > 0.20$) (31). Considering the unknown factorial structure of the responses to the FIQR, an estimate of the number of relevant factors was obtained by means of parallel analysis (PA) (32), and then an exploratory factor analysis (EFA) for ordinal data (33) with oblique (Promax) rotation on a randomly split half of the dataset ($n=242$) was used to study the contribution of each item to the factors identified by the PA. An item was considered as contributing to a factor when factor loading was 0.4 or more: this threshold was chosen taking into account sample size, number of items and the exploratory nature of the study (34). As mono-dimensionality was not confirmed, the number of underlying factors and their relation to each item was used to separate the scale into subscales, the clinical acceptability of which was evaluated by experts. Each subscale underwent further analyses after verifying a good fit between the data and the model using confirmatory factor analysis (CFA) on the second half of the dataset ($n=243$). The following indices were used to evaluate the fit of the data to the model: the Tucker-Lewis fit index (TLI), the comparative fit index (CFI), the root mean square error of approximation (RMSEA), and the standardised root mean square residual (SRMR). For an acceptable fit, TLI and CFI should be ≥ 0.95 , RMSEA ≤ 0.08 and SRMR ≤ 0.10 (35). Once the factorial structure of the scale was determined, the matrix of the individual raw scores of each subject ($n=503$) on each of the two selected subscales underwent Rasch analysis using WINSTEPS software v. 3.68.2 (36). The technical aspects of the analysis have been described elsewhere (37, 38). As a first step, we used a rating scale

Table I. Scores obtained on the FIQR and FAS (mean and median values, standard deviations and 95% confidence intervals for each item, sub-dimensions and total score of the FIQR, and for total score of the FAS) by the study patients ($n=503$).

| Item n | Item description | Mean | Median | Standard deviation | 95% confidence interval | Score range |
|--------|--|--------------|--------------|--------------------|-------------------------|--------------|
| 1 | Brush or comb hair | 2.48 | 1.00 | 2.77 | 2.24–2.72 | 0–10 |
| 2 | Walk continuously for 20 minutes | 4.54 | 5.00 | 3.00 | 4.27–4.80 | 0–10 |
| 3 | Prepare a homemade meal | 3.28 | 3.00 | 2.76 | 3.03–3.52 | 0–10 |
| 4 | Vacuum, scrub or sweep floors | 4.67 | 5.00 | 2.98 | 4.41–4.93 | 0–10 |
| 5 | Lift and carry a bag full of groceries | 5.32 | 5.00 | 2.83 | 5.07–5.57 | 0–10 |
| 6 | Climb one flight of stairs | 4.33 | 5.00 | 3.09 | 4.06–4.61 | 0–10 |
| 7 | Change bed sheets | 4.04 | 4.00 | 2.92 | 3.78–4.30 | 0–10 |
| 8 | Sit in a chair for 45 minutes | 4.49 | 5.00 | 3.08 | 4.21–4.76 | 0–10 |
| 9 | Go shopping for groceries | 3.86 | 4.00 | 2.88 | 3.61–4.12 | 0–10 |
| | <i>FIQR function sub-total</i> | <i>12.32</i> | <i>12.00</i> | <i>6.90</i> | <i>11.72–12.93</i> | <i>0–30</i> |
| 10 | Cannot achieve goals | 4.72 | 5.00 | 2.76 | 4.48–4.96 | 0–10 |
| 11 | Feel overwhelmed | 3.26 | 3.00 | 3.15 | 2.98–3.54 | 0–10 |
| | <i>FIQR overall impact sub-total</i> | <i>7.98</i> | <i>7.00</i> | <i>5.24</i> | <i>5.52–8.44</i> | <i>0–20</i> |
| 12 | Pain rating | 5.94 | 6.00 | 2.10 | 5.76–6.13 | 0–10 |
| 13 | Fatigue rating | 6.42 | 7.00 | 2.36 | 6.21–6.63 | 0–10 |
| 14 | Stiffness rating | 5.76 | 6.00 | 2.21 | 5.57–5.96 | 0–10 |
| 15 | Sleep quality | 5.82 | 6.00 | 2.63 | 5.59–6.05 | 0–10 |
| 16 | Depression level | 4.49 | 5.00 | 2.46 | 4.27–4.70 | 0–10 |
| 17 | Memory problems | 5.25 | 5.00 | 2.58 | 5.02–5.48 | 0–10 |
| 18 | Anxiety level | 5.32 | 6.00 | 2.58 | 5.09–5.54 | 0–10 |
| 19 | Tenderness level | 4.90 | 5.00 | 2.50 | 4.68–5.12 | 0–10 |
| 20 | Balance problems | 4.22 | 4.00 | 2.39 | 4.01–4.43 | 0–10 |
| 21 | Environmental sensitivity | 5.72 | 6.00 | 2.56 | 5.49–5.94 | 0–10 |
| | <i>FIQR symptoms sub-total</i> | <i>27.17</i> | <i>27.00</i> | <i>8.29</i> | <i>26.45–27.90</i> | <i>0–50</i> |
| | <i>FIQR TOTAL SCORE</i> | <i>47.32</i> | <i>47.00</i> | <i>18.09</i> | <i>45.73–48.90</i> | <i>0–100</i> |
| | <i>FAS TOTAL SCORE</i> | <i>5.74</i> | <i>5.71</i> | <i>1.71</i> | <i>5.59–5.89</i> | <i>0–10</i> |

FIQR: Fibromyalgia Impact Questionnaire Revised version; FAS: Fibromyalgia Assessment Status.

model to investigate whether the rating scale of the FIQR was being used in the expected manner. The response categories were evaluated and collapsed using standardised procedures (39). After rating scale modifications, a new series of Rasch analyses of the two subscales were performed as suggested by the preliminary dimensionality analysis. The validity of each scale was assessed by evaluating the fit of the individual items to the latent trait in accordance with the Rasch model and examining whether the pattern of item difficulties was consistent with the model expectations. Information-weighted (infit) and outlier-sensitive (outfit) mean-square statistics (MnSq) for each item were calculated, considering a MnSq value between 0.8 and 1.2 as indicating an acceptable fit (37). Reliability was evaluated in terms of 'separation', which was defined as the ratio between the true spread of the measures and their measurement er-

ror. A separation of 2.0 is considered good, and a related index is the reliability of these separation indices, which provides the degree of confidence that can be placed in the consistency of the estimates (range 0–1; coefficients of >0.80 are considered good, and >0.90 excellent) (37). A principal component analysis (PCA) of the standardised residuals was used to investigate the local independence of the items (36) and the presence of sub-dimensions as an independent confirmation of the mono-dimensionality of the scale. The following criteria were used to confirm mono-dimensionality: a) a cut-off of 50% of the variance explained by the trait that the scale is intended to measure (the 'Rasch factor'); and b) the eigenvalue of the first residual factor smaller than three (36).

In addition, convergent validity of the two FIQR subscales ("function"; "symptoms") was assessed by calculat-

ing the Pearson correlation coefficient (r) between their raw scores and that of the FAS. STATA 10.1 (StataCorp, LP College Station, TX, USA) was used to perform the PA. Lisrel 8.80 (Scientific Software International, Inc. Lincolnwood, IL, USA) was used for the EFA and CFA.

Results

Descriptive statistics

Table I shows the score details for each item, subscales and total score of the FIQR, as well as for the FAS. Frequency distribution of FIQR scores was symmetrical, with a positive excess kurtosis (mean 47.32; median 47; skewness 0.20; kurtosis 2.54). The spidergram in Figure 2 shows the distribution of the FIQR scores in the patient sample as a whole: it offers a simplified means to visualise complex results across all domains of FIQR in a single figure. The five highest scoring items (greater disease impact) were related to symptoms: fatigue/energy (FIQR13), pain (FIQR12), stiffness (FIQR14), sleep quality (FIQR15), and environmental sensitivity (FIQR21). The lowest scoring items included functional activities, such as brushing/combing hair, preparing a home-made meal, shopping for groceries, and changing bed sheets.

Classical test theory analysis

The Cronbach alpha of the FIQR was 0.94. All items had an item-to-total correlation between 0.41 (item 21) and 0.78 (item 9). After randomly splitting the dataset into two halves, PA applied to the first half revealed two factors with empirical eigenvalues that exceeded those averaged from 200 random datasets with the same number of variables and observations, that respectively explained 57.9% and 8.6% of the variance (66.5% cumulatively). In the subsequent two factor EFA, results of which are presented in Table II, items 10 and 11, supposed to represent “overall impact”, failed to achieve salient loadings.

On the basis of these findings (consistent with expert opinion), two separate subscales were identified: the first including FIQR items 1–9 (“function”), the second composed of items 12–21

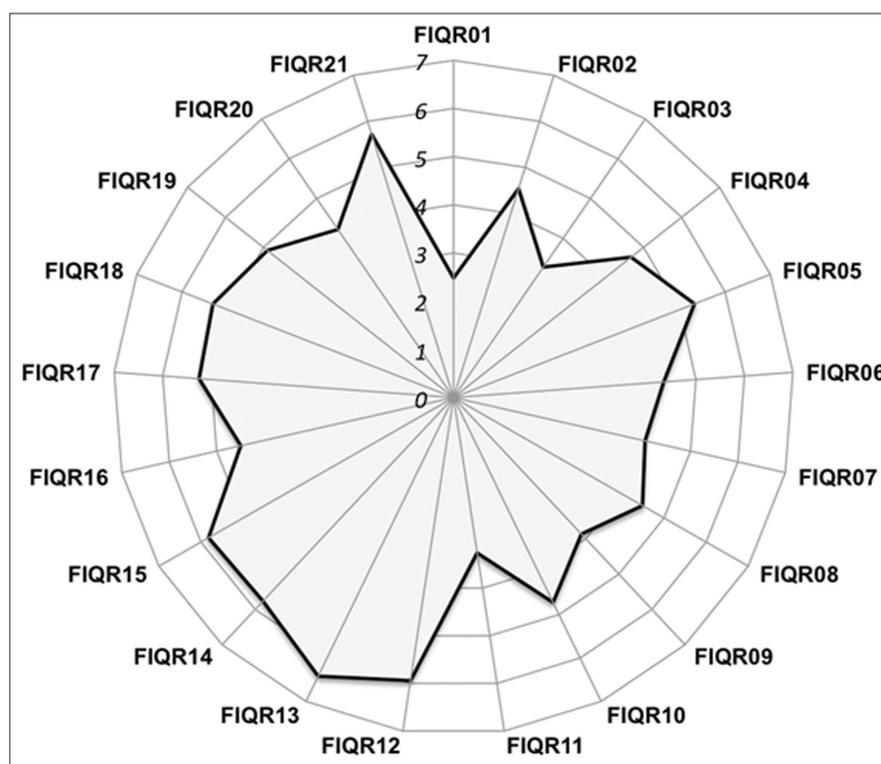


Fig. 2. Spidergram of the FIQR domains of the overall cohort of fibromyalgia patients ($n=503$). The domain scores are plotted from 0 (worst, at the centre) to 10 (best, at the outside). For item numbers, see Tables I and II.

(“symptoms”). Items 10 and 11 were therefore not analysed further.

A CFA of the 2-factor model on the second half of the dataset showed a borderline fit: TLI, CFI, RMSEA, and SRMR were respectively 0.94, 0.95, 0.11, and 0.068.

As expected, a good correlation was found between the FAS total score and both FIQR subscales ($r=0.64$ for “function”; $r=0.84$ for “symptoms”). The correlation between FIQR and FAS was good ($r=0.78$). The correlations between FIQR and FAS subdomains (pain, fatigue and sleep) were moderate and good ($r=0.46$, $r=0.65$ and $r=0.61$, respectively).

Rasch analysis

Rating scale diagnostics showed that the rating categories did not comply with the pre-set criteria for category functioning (average measures, thresholds, etc.), suggesting that raters had difficulty in discerning the eleven response levels proposed by the original FIQR. Accordingly, the response categories were revised: the pre-set criteria were best met by collapsing categories

1–3, 4–6, and 7–9 (Fig. 3). As shown in Table II, after combining these rating scale categories, all the items of each of the subscales (“function” and “symptoms”) fitted their respective constructs (MnSq 0.8–1.2) except: a) the underfitting item 1 “Brush or comb hair” in the function subscale (Infit Mnsq=1.44; Outfit Mnsq=1.29); b) the borderline items 17 “Memory problems” (Infit Mnsq=1.21; Outfit Mnsq=1.23) and 21 “Sensitivity to loud noises, bright lights, odours and cold” (Infit Mnsq=1.21; Outfit Mnsq=1.13) in the symptom subscale; and c) the slightly overfitting item 3 “Prepare a homemade meal” (Infit Mnsq=0.78; Outfit Mnsq=0.76) and item 9 “Go shopping for groceries” (Infit Mnsq=0.75; Outfit Mnsq=0.70) in the function subscale, and item 14 “Level of stiffness” in the symptom subscale (Infit Mnsq=0.76; Outfit Mnsq=0.75). Table III summarises the results for subject ability, item difficulty, reliability indices, and the two PCAs of the standardised residuals. The sample-item matching was very good for “Symptoms” (average subject ability=0.03) and acceptable for “Function” (average sub-

Table II. Exploratory factor analysis of a two-factor solution as suggested by the parallel analysis (loadings of >0.40 are in bold) (sample size: n=242) and Rasch analysis with item calibration and fit information (under-fitting values in bold and over-fitting values in italics; sample size: n=503). Items with higher (positive) Measure in Rasch analysis are those showing lower level of difficulty.

| Item | Factor analysis | | Rasch analysis | | |
|--|-----------------|--------------|----------------|-------------|-------------|
| | Factor 1 | Factor 2 | Measure (SE) | Infit MnSq | Outfit MnSq |
| 1 Brush or comb hair | 0.638 | -0.029 | 1.20 (0.07) | 1.44 | 1.29 |
| 2 Walk continuously for 20 minutes | 0.680 | 0.168 | -0.31 (0.06) | 0.86 | 0.83 |
| 3 Prepare a homemade meal | 0.894 | -0.095 | 0.61 (0.06) | <i>0.78</i> | <i>0.76</i> |
| 4 Vacuum, scrub or sweep floors | 0.722 | 0.011 | -0.41 (0.06) | 1.16 | 1.11 |
| 5 Lift and carry a bag full of groceries | 0.665 | 0.134 | -0.88 (0.06) | 0.81 | 0.79 |
| 6 Climb one flight of stairs | 0.460 | 0.387 | -0.16 (0.06) | 1.11 | 1.07 |
| 7 Change bed sheets | 0.737 | 0.031 | 0.03 (0.06) | 0.96 | 0.90 |
| 8 Sit in a chair for 45 minutes | 0.823 | -0.125 | -0.27 (0.06) | 1.15 | 1.19 |
| 9 Go shopping for groceries | 0.701 | 0.174 | 0.19 (0.06) | <i>0.75</i> | <i>0.70</i> |
| 10 Cannot achieve goals | 0.395 | 0.378 | – | – | – |
| 11 Feel overwhelmed | 0.288 | 0.343 | – | – | – |
| 12 Pain rating | 0.140 | 0.583 | -0.48 (0.07) | 0.87 | 0.86 |
| 13 Fatigue rating | -0.026 | 0.736 | -0.80 (0.07) | 1.02 | 1.00 |
| 14 Stiffness rating | 0.146 | 0.629 | -0.24 (0.07) | <i>0.76</i> | <i>0.75</i> |
| 15 Sleep quality | -0.088 | 0.739 | -0.30 (0.07) | 1.03 | 1.05 |
| 16 Depression level | 0.295 | 0.417 | 0.66 (0.06) | 1.03 | 1.07 |
| 17 Memory problems | -0.029 | 0.530 | 0.06 (0.07) | 1.21 | 1.23 |
| 18 Anxiety level | 0.011 | 0.668 | 0.05 (0.07) | 1.07 | 1.03 |
| 19 Tenderness level | 0.243 | 0.409 | 0.39 (0.06) | 1.02 | 1.09 |
| 20 Balance problems | 0.213 | 0.541 | 0.83 (0.06) | 0.80 | 0.81 |
| 21 Environmental sensitivity | -0.111 | 0.598 | -0.19 (0.07) | 1.21 | 1.13 |

ject ability=-0.90). Separation indexes and respective reliability were satisfactory for both subscales (person separation reliabilities ≥ 0.83 ; item separation reliabilities ≥ 0.98). Subject ability span was good (>11 logits in each subscale), whereas item difficulty span ranged from 1.65 logits (symptoms) to 2.08 logits (function). The two PCAs on the residuals did not sustain the presence of additional dimensions, thus confirming the unidimensionality of each subscale, while the absence of residual correlations >0.30 between item pairs verified the local independence of their items.

Discussion

Patient-reported outcome measures such as the FIQR are increasingly being used as a key part in the evaluation of clinical health care interventions, but their ability to improve decision making greatly relies on their psychometric strength (18). Bennet *et al.* proposed the FIQR in 2009 as an extensive modification of the previous questionnaire in terms of content, number of items and response options (20). It was designed

to overcome some of the drawbacks with the FIQ (22), but its validation (and that of the subsequent translations and cultural adaptations) used only a classical test theory approach.

New and more detailed psychometric approaches, including a mix of classical test theory and item response theory methods, have recently been recommended to provide reliability and validity estimates of the instrument (40). This study is the first psychometric analysis of FIQR using both CTT and Rasch analysis to examine its properties in depth and provide a rationale for improving its metric quality. As such we mainly concentrated on analysing its dimensions and rating scale diagnostics, and identifying the appropriateness of its items for measuring the intended construct.

The five most endorsed (*i.e.* highest scoring) items are all related to symptoms: fatigue/energy, pain, stiffness, sleep quality, and environmental sensitivity. These findings are very similar to those reported by Bennett *et al.* (20) and Ediz *et al.* (25), thus supporting the im-

portance of these domains for a patient-centred evaluation (15, 30, 41).

Cronbach's alpha and the item-to-total statistics showed acceptable values according to CTT rules. This finding is similar to that of Bennett *et al.* (Cronbach's alpha 0.95) (20), and more favourable than that of Ediz *et al.* (25) and Srifi *et al.* (24), who respectively reported alpha values of 0.89–0.91 and 0.91–0.92. However, in terms of mono-dimensionality, the factor analysis (parallel analysis and EFA) hinted at two factors: the first factor consisting of the first nine items (related to the “function” domain), the second consisting of the last 10 items (no. 12–21) related to the “symptoms” domain. These are two of the three domains defined by the original authors in creating the scale. The third domain – “overall impact” (consisting of questions no. 10 and 11) – did not emerge as an independent factor, which is in line with the observation that, under the best conditions, a minimum of three items is critical for producing a dimension in factor analysis (42). These three domains are often used separately to validate the scale (20, 24, 25) and, in terms of total score, the “function” domain accounts for 30%, “symptoms” for 50%, and “overall impact” for 20% (20). This weighting system seems to be based on theoretical considerations (expert evaluation resulting in different weighting of the three domains), rather than a statistical approach, and we think it warrants further research.

Rasch analysis of the 11 rating scale categories showed that the FIQR has many disordered thresholds. The category thresholds are the ability levels at which the response to either of the two adjacent categories is equally likely. Disordered thresholds occur when respondents have difficulty consistently discriminating between response options (*e.g.* because there are too many response options or the labelling of the options is confusing). Our findings indicate that our subjects were able to discern only five categories (rather than the original eleven levels), and hence it might be appropriate to simplify the item format.

This is not the first time that an 11-level numeric rating scale (0–10) has failed

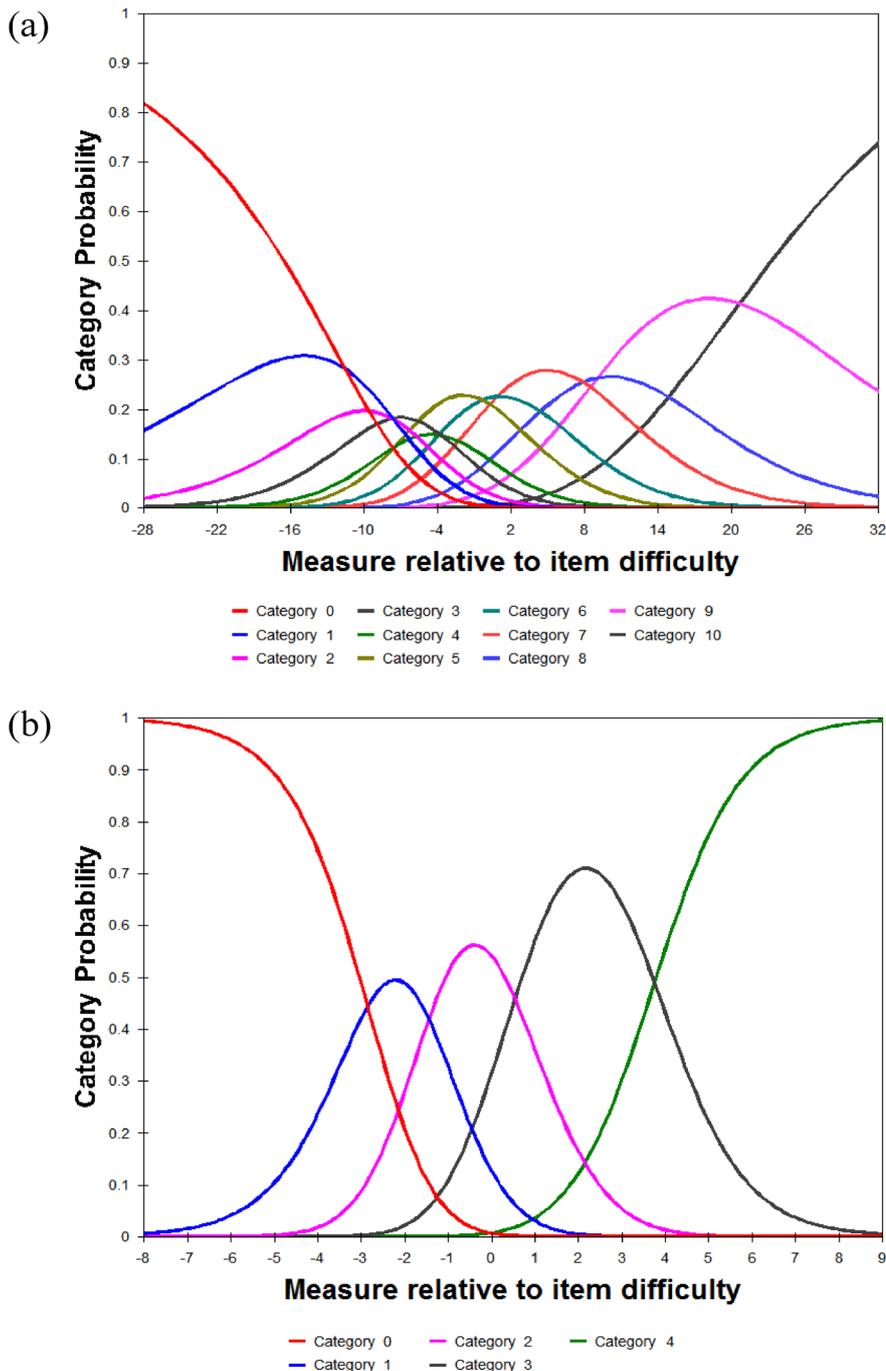


Fig. 3. Category probability curves: **a)** Curves of the original 11 categories (0–10); and **b)** of the 5 revised categories obtained by combining the original categories as follows: 0=0; 1–3=1; 4–6=2; 7–9=3; 10=4. The y axis represents the probability (0–1) of responding to one of the rating categories and the x axis the different performance values (patient ability minus item difficulty) in logits. The ideal plot should look like an ordered even succession of hills, with an ‘emerging’ crest where each category is modal over a certain range. The “0” curve declines as the subject’s ability increases; the crossing point (where 0 and 1 are equally probable) is the first “threshold”. The same applies to the other curves. The graph in (a) shows that the probability of using some categories is never higher than that of other adjacent ratings, whereas the graph in (b) shows that the probability of selecting each of the five revised rating categories is a clear function of the level of ability shown by the subject in the x axis.

to demonstrate appropriate functioning. For example, an analysis of the mobility section of the Prosthesis Evaluation Questionnaire (a self-administered

questionnaire designed to evaluate subjects who have undergone amputation) showed that patients were unable to distinguish their abilities as finely as

suggested by the 0–10 numeric rating scale, whereas five well-selected response categories were able to improve the measurement qualities of the scale without decreasing its reliability indices (43). Similarly, when Khadka *et al.* (44) used Rasch analysis to assess the performance of the rating scales of seventeen outcome measures in ophthalmology, they found that a simple and uniform question format (*e.g.* four or five well-labelled categories) was more likely to function (from a psychometric point of view) than those with a complicated question format and a large number of response categories.

As for the fit of the individual items to the latent trait (analysis of internal construct validity), the item “Brush or comb hair” (that shows the lower level of difficulty in the function subscale) in the function subscale was underfitting, which indicates either that it does not tap the same underlying construct as the other items, or that it is poorly written or, more probably, too sensitive to (personal and environmental) confounding factors. This item was added to the FIQR without any formal validation at item level, following a suggestion from the focus group. If our findings are confirmed, it might be opportune to consider replacing it with one that is more homogeneous with the scale’s conceptual framework, and capable of optimising content validity, coverage and the technical quality of the measure in line with modern psychometric approaches. Moreover, the items “Memory problems” and “Sensitivity to loud noises, bright lights, odours and cold” (in the “symptoms” subscale) had a borderline fit. They warrant further investigation but, for the present, we suggest retaining both as the misfit is minor and they are clinically important. Similarly, the items “Prepare a homemade meal” and “Go shopping for groceries” in the “function” subscale and “Level of stiffness” in the “symptoms” subscale present a too predictable pattern at Rasch analysis (overfit) and could perhaps be deleted (as redundant) or substituted. Alternatively, they could be retained if considered clinically relevant because overfitting simply tends to overestimate differences in raw scores,

Table III. Subjects' ability levels and reliability indices for the two FIQR subscales (n=503).

| | Subscale 1 "Function" (items 1–9) | | Subscale 2 "Symptoms" (items 12–21) | |
|--|--------------------------------------|------------------------------------|--|--------------------------------------|
| | Original coding | Recoded | Original coding | Recoded |
| Average ability levels (range) | -0.41 (-3.79–2.92) | -0.90 (-6.08–5.62) | 0.03 (-4.16–2.73) | 0.03 (-6.51–5.30) |
| SD | 0.91 | 1.80 | 0.60 | 1.36 |
| Person separation index (person separation reliability) | 2.26 (0.84) | 2.70 (0.88) | 2.39 (0.85) | 2.23 (0.83) |
| Cronbach's alpha | 0.92 | 0.91 | 0.87 | 0.85 |
| Range of item difficulty estimates | -0.38 (item 5) to 0.53 (item 1) | -0.88 (item 5) to 1.20 (item 1) | -0.34 (item 13) to 0.36 (item 20) | -0.80 (item 13) to 0.85 (item 20) |
| SD | 0.25 | 0.57 | 0.20 | 0.49 |
| Item separation index (item separation reliability) | 9.38 (0.99) | 8.55 (0.99) | 7.70 (0.98) | 6.99 (0.98) |
| Variance explained by the Rasch factor | 62.9% | 58.4% | 49.0% | 44.8% |
| Eigenvalue of the first residual factor | 2.2 | 2.2 | 1.8 | 1.7 |
| Variance explained by the first residual factor | 9.0% | 10.2% | 9.3% | 9.4% |

without affecting dimensionality or severely degrading the metric properties of the scale.

The item difficulty range of both subscales (Table III) covers an acceptable distance of each targeted construct (between 1.6 and 2 logits). The targeting of item difficulty to patient ability (*i.e.* the extent to which the items are appropriately difficult for the sample) was very good in the case of 'symptoms' and acceptable in the case of 'function'. However, a study of the stability of item hierarchy (differential item functioning) across sub-samples defined on the basis of potentially relevant clinical criteria (*e.g.* age, gender or symptom patterns) is warranted.

A number of potential limitations of this study should be mentioned. No additional dimension was suggested by RA, but the CFA of the 2-factor model on the second half of the dataset showed only a borderline fit that could be due to the limited sample size. We cannot exclude that some specific (linguistic, cultural or technical) characteristics of the Italian version of the FIQR may have influenced some results, although this version was obtained by means of a thorough forward/backward translation process. A further caveat is that this analysis was based on a population of adults from a relatively limited geographical area of Italy. Finally, the present findings need to be confirmed

in different countries, contexts and patient samples.

Conclusion

To our knowledge, this is the first study investigating the dimensionality and validity of the FIQR using stringent criteria that included item response theory methods.

Our main findings are:

- two of the three domains were confirmed, whereas the third probably needs some more items in order to become an interpretable domain from a psychometric point of view;
- the selection of the items in the two analysed domains seems to be appropriate (also in terms of coverage and technical quality), although there is room for some refinement;
- the 0–10 rating scale should be simplified. A five-level response structure seems to be enough, but its actual performance needs to be analysed. This preliminary study supports the usefulness of this questionnaire for measuring the multidimensional spectrum of FM-related functional (45–47) and HRQL problems, and provides a basis for further research aimed at improving its measurement qualities.

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